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Comparing multiple treatments 2: statistical methods for network meta-analysis

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A plethora **meta-analyses** has been published in the last years

"Although **Mirtazapine** is likely to have a faster onset of action than **Sertraline and Paroxetine** no significant differences were observed..."

"...statistically significant differences in terms of efficacy between **Fluoxetine and Venlafaxine**, but the clinical meaning of these differences is uncertain..."

"...**meta-analysis** highlighted a trend in favour of **Sertraline** over other **Fluoxetine**"

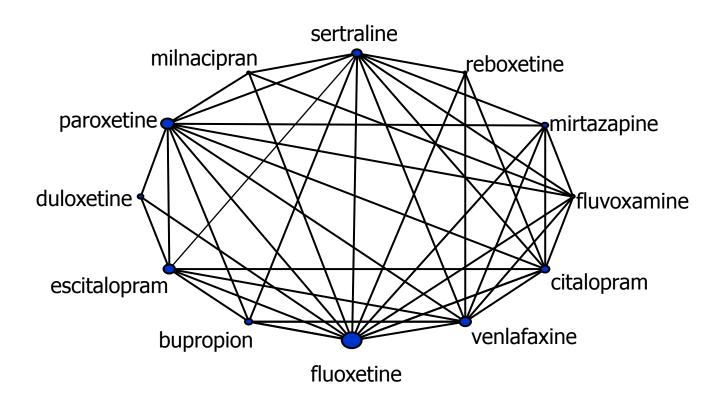
"**Venlafaxine** tends to have a favorable trend in response rates compared with **duloxetine**"

Fluoxetine: 28€

Venlafaxine:111€

Sertaline: 76 €





Comparative efficacy and acceptability of 12 new-generation antidepressants: a multiple-treatments meta-analysis

Andrea Cipriani, Toshiaki A Furukawa, Georgia Salanti, John R Geddes, Julian P T Higgins, Rachel Churchill, Norio Watanabe, Atsuo Nakagawa, Ichiro M Omori, Huqh McGuire, Michele Tansella, Corrado Barbui

Summary

Lancet 2009; 373: 746-58

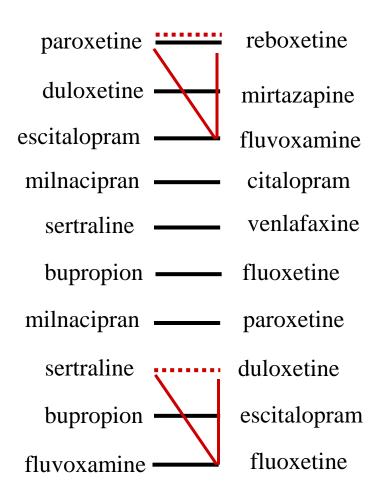
Published Online January 29, 2009 DOI:10.1016/50140-6736(09)60046-5

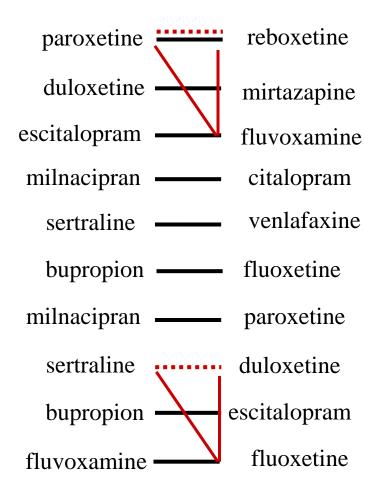
See Comment page 700

Department of Medicine and Public Health, Section of Boychister, and Clinical **Background** Conventional meta-analyses have shown inconsistent results for efficacy of second-generation antidepressants. We therefore did a multiple-treatments meta-analysis, which accounts for both direct and indirect comparisons, to assess the effects of 12 new-generation antidepressants on major depression.

Methods We systematically reviewed 117 randomised controlled trials (25928 participants) from 1991 up to Nov 30, 2007, which compared any of the following antidepressants at therapeutic dose range for the acute treatment of unipolar major depression in adults: bupropion, citalopram, duloxetine, escitalopram, fluoxetine, fluoxamine, milnacipram mirtazapine paroxetine reboxetine sertraline and venlafaxine. The main outcomes were the proportion

paroxetine ——	reboxetine	
duloxetine ——	mirtazapine	
escitalopram ——	fluvoxamine	
milnacipran ——	citalopram	
sertraline ——	venlafaxine	
bupropion ——	fluoxetine	
milnacipran ———	paroxetine	
sertraline ?	duloxetine	
bupropion ——	escitalopram	
fluvoxamine ———	fluoxetine	





paroxetine 0% 7% sertraline 0% citalopram escitalopram 26% fluoxetine 0% fluvoxamine 0% milnacipran 1% venlafaxine 11% 0% reboxetine bupropion 0% mirtazapine 54% duloxetine 0%

Probability of being the best

4 Fluoride modalities for preventing dental carries: series of pairwise metaanalyses

Treatment comparison		Studies
Placebo	Toothpaste	69
	Gel	13
	Rinse	31
	Varnish	3
Toothpaste	Rinse	6
Toothpaste	Varnish	1
Gel	Rinse	1
Gel	Varnish	?

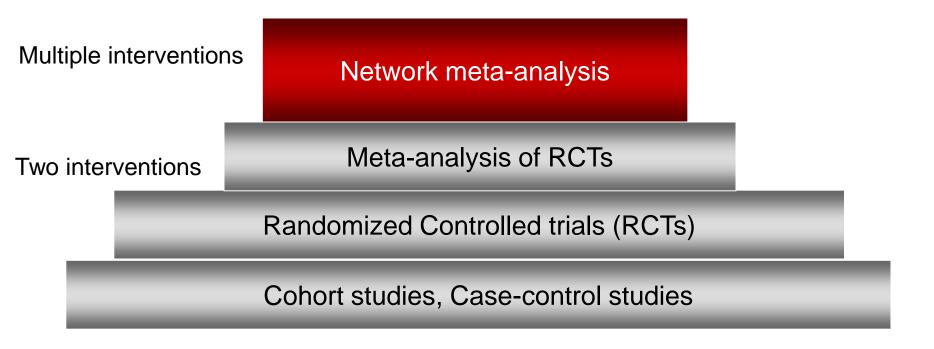
Multiple treatments and series of meta-analyses

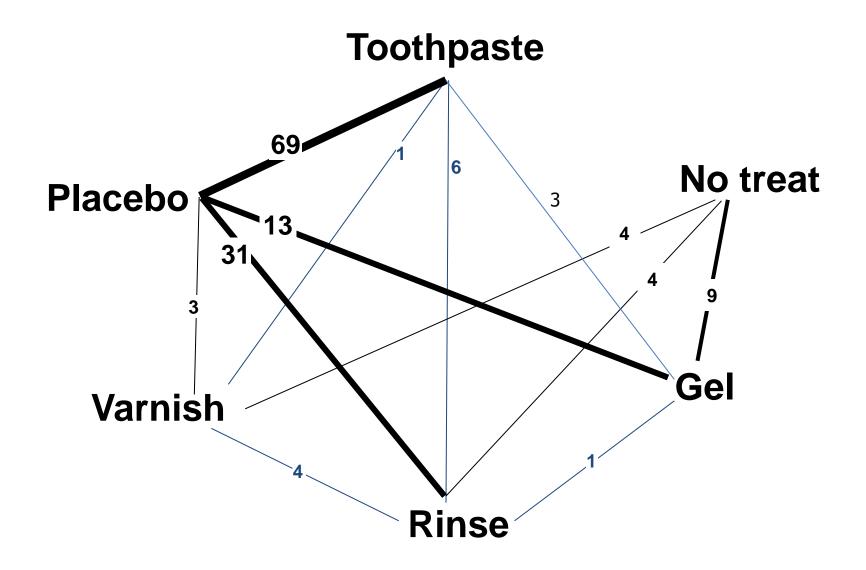
With pairwise meta-analyses we cannot answer the following questions:

- Which fluoride modality is the best?
- What is the ranking of fluoride treatments according to effectiveness?
- Which is better: Gel or Varnish (0 studies)

A new methodological framework

Other names: Multiple treatments meta-analysis, Mixed treatment comparisons

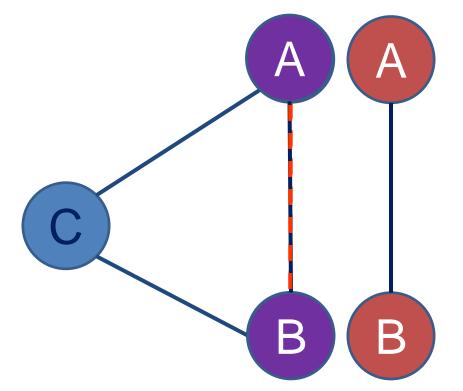




Aims of the workshop

- To explain indirect and mixed comparison of interventions
 - Assumptions
 - Statistical methods
- To understand the statistical models for network meta-analysis
- To discuss presentation of results from network meta-analysis
- To understand inconsistency models

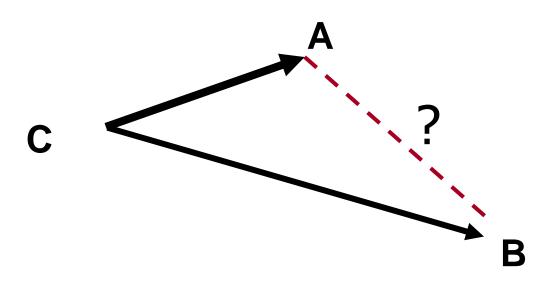
Indirect and mixed effects



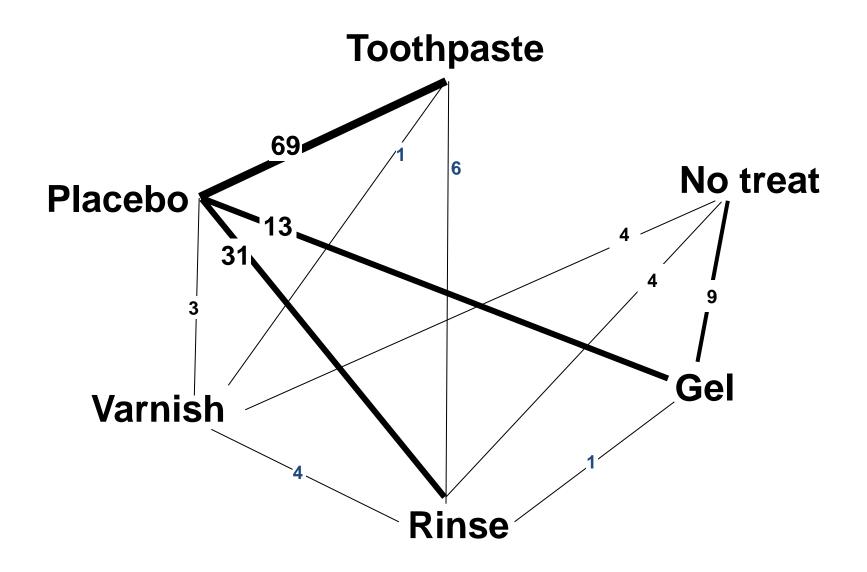
Indirect effect Direct effect Mixed effect

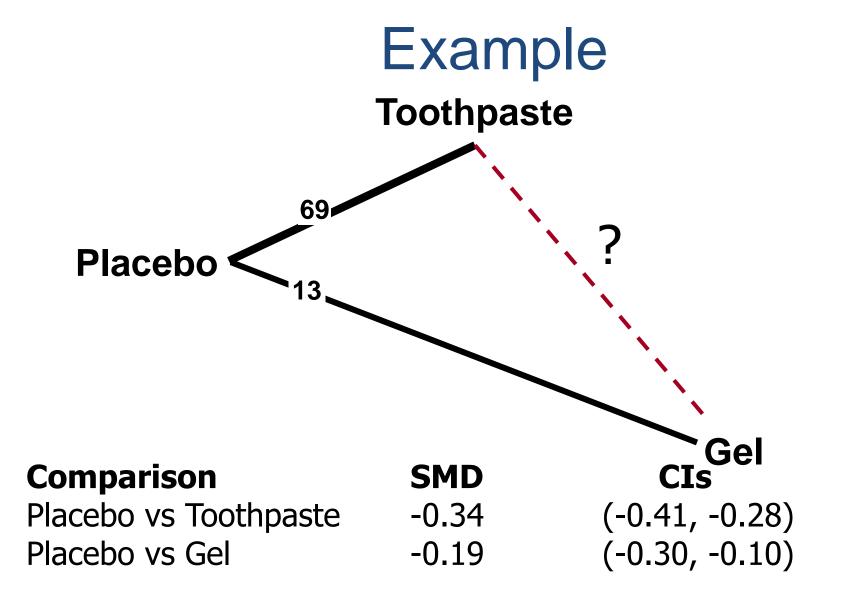
Indirect comparison

• We can obtain an *indirect* estimate for A vs B from RCTs comparing A vs C and B vs C:



 $SMD_{AB} = SMD_{AC} + SMD_{CB}$ $SMD_{AB} = SMD_{AC} - SMD_{BC}$ $Var(SMD_{AB}) = Var(SMD_{AC}) + Var(SMD_{BC})$

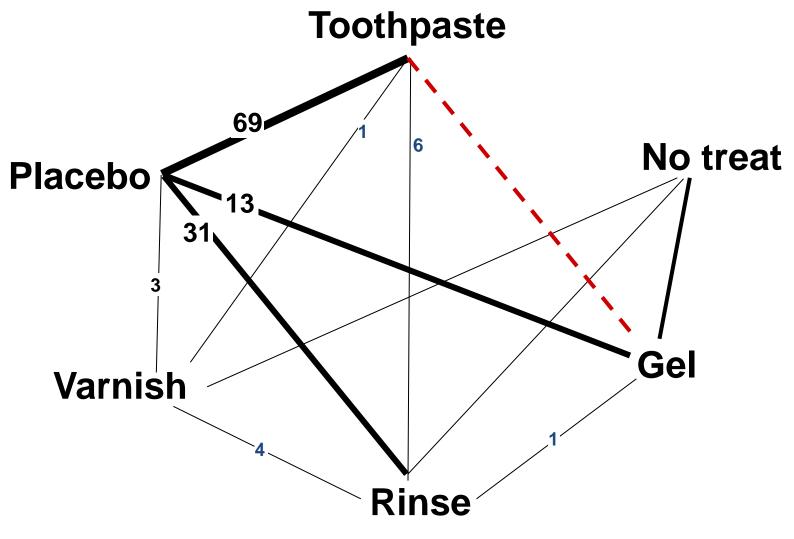




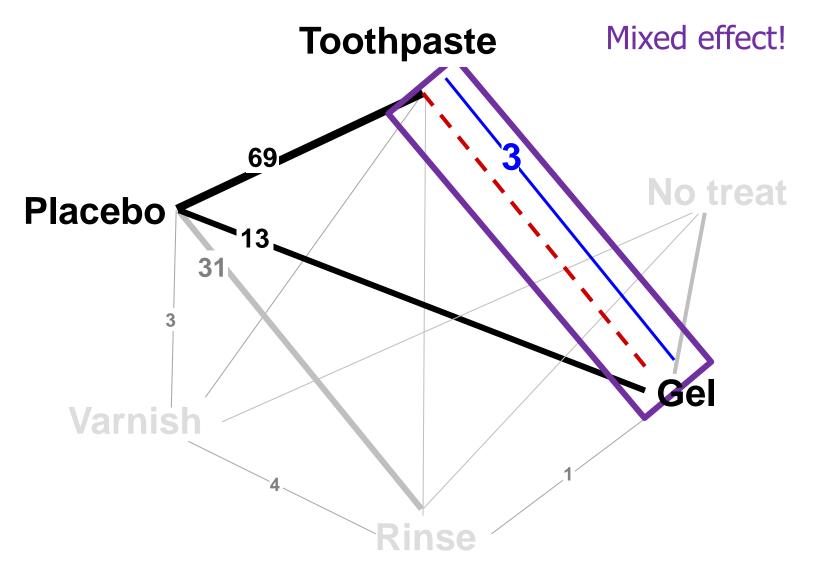
How to compare Gel to Toothpaste?

Exercise

Indirect SMD_{GvsT} = SMD_{PvsT} - SMD_{PvsG} Indirect $SMD_{GvsT} = -0.34 - (-0.19) = -0.15$ Variance Indirect SMD_{GvsT} = Variance SMD_{PvsT} + Variance SMD_{PvsG} Variance $SMD_{PVST} = ((high CI - low CI)/3.92)^2$ Variance $SMD_{PvsT} = ((-0.28 - (-0.41))/3.92)^2 = 0.0011$ Variance $SMD_{GVST} = ((-0.10 - (-0.30))/(3.92)^2) = 0.0026$ Variance Indirect $SMD_{GvsT} = 0.0011 + 0.0026 = 0.0037$ **SE Indirect SMD**_{GvsT} = sqrt(0.0037) = 0.061 **95% CI for Indirect** $SMD_{GvsT} = (-0.15 - 1.96 \times 0.061, -0.15 + 1.96 \times 0.061)$ **95% CI for Indirect** $SMD_{GvsT} = (-0.27, -0.03)$



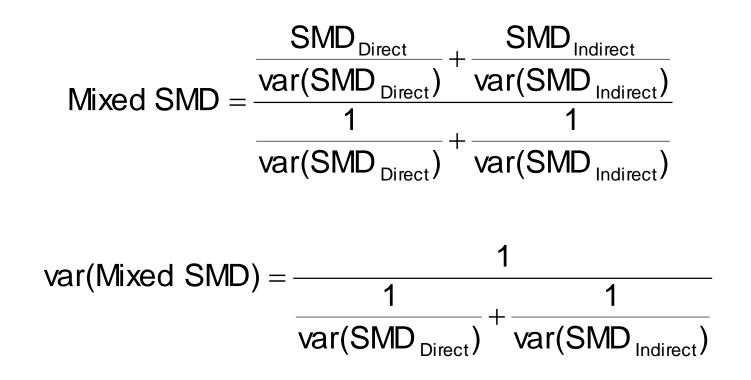
Indirect SMD Gel vs Toothpaste: -0.15 (0.0037)



Indirect SMD Gel vs Toothpaste: -0.15 (0.0037)

Mixed comparison

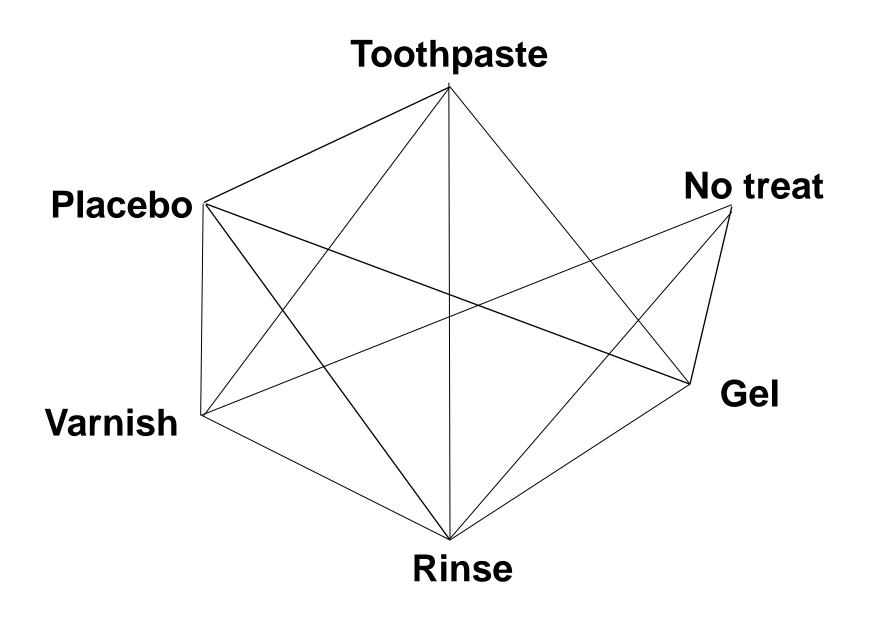
 Summarize direct and indirect effect size into a single mixed effect



Mixed comparison

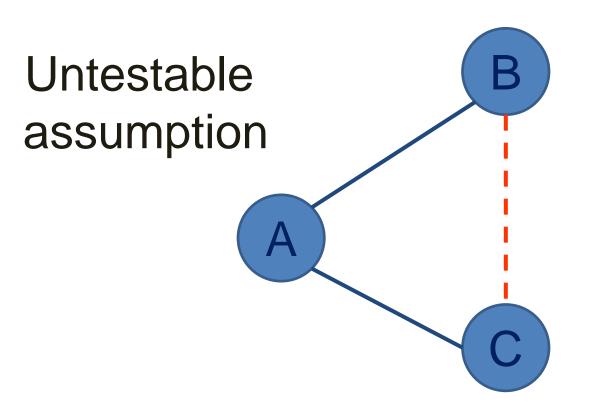
Indirect $SMD_{GvsT} = -0.15$ $Var(Indirect <math>SMD_{GvsT}) = 0.004$ Direct $SMD_{GvsT} = 0.04$ $Var(Direct <math>SMD_{GvsT}) = 0.011$ $Var(Direct SMD_{GvsT}) = 0.011$ $Var(Direct SMD_{GvsT}) = 0.011$

You can do this with any measure... InOR, InRR, RD, mean difference, HR, Peto's InOR etc...



Extend the idea of mixed effect sizes in the entire network

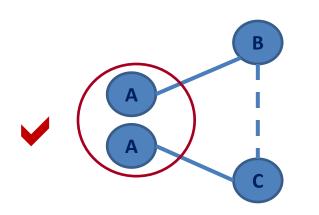
Transitivity



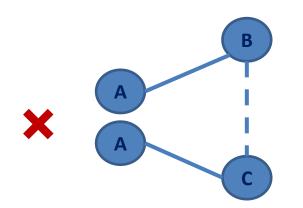
The anchor treatment A is 'transitive'

....but you can evaluate clinically and epidemiologically its plausibility

Transitivity requires... (1)



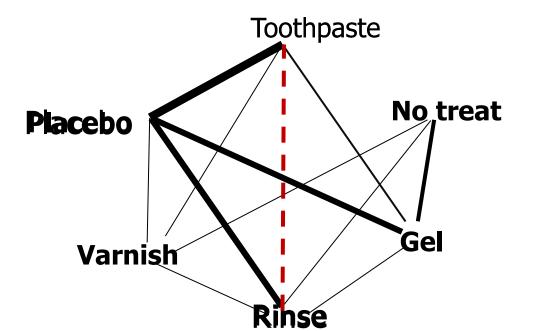
The 'anchor' treatment A to be similarly defined when it appears in AB and AC trials. e.g. a treatment given at different doses but no systematic difference in the *average* dose of A across AB and AC studies



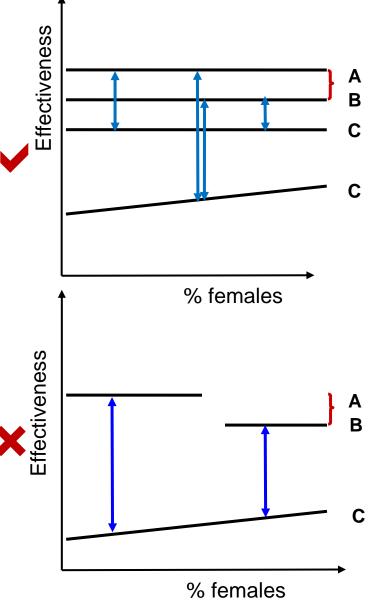
The 'anchor' treatment A may be different in AB and AC studies e.g. injection versus pill

Transitivity requires... (1)

- However, placebo toothpaste and placebo rinse might not be comparable as the mechanical function of brushing might have a different effect on the prevention of caries.
- If this is the case, the transitivity assumption is doubtful (Salanti 2009).



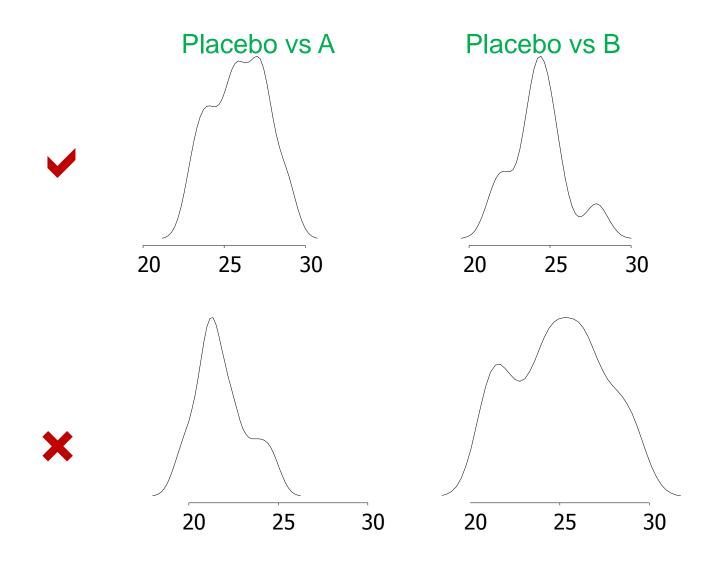
Transitivity means that...



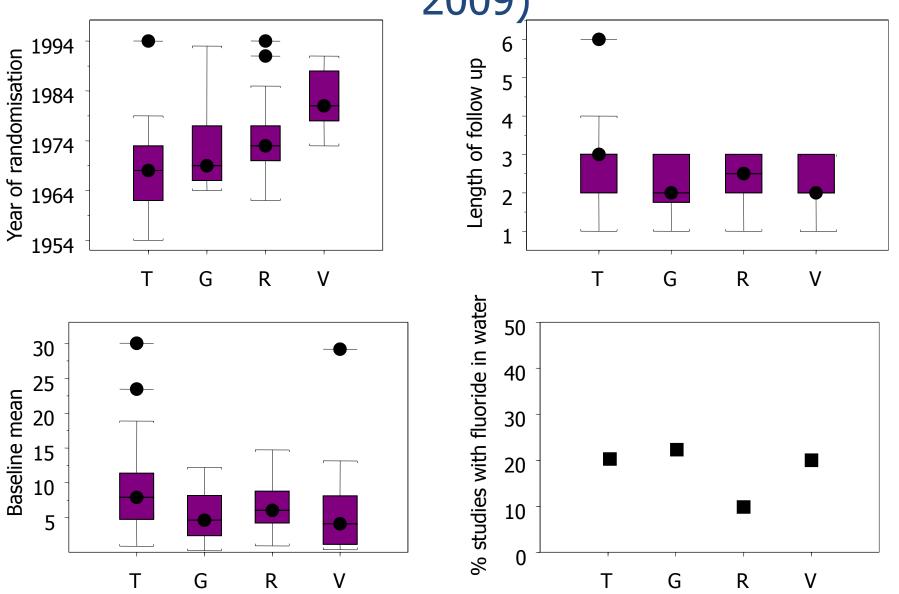
AC and BC trials do not differ with respect to the distribution of effect modifiers

Difficult to defend when you have older and newer treatments, and variables are often unobserved

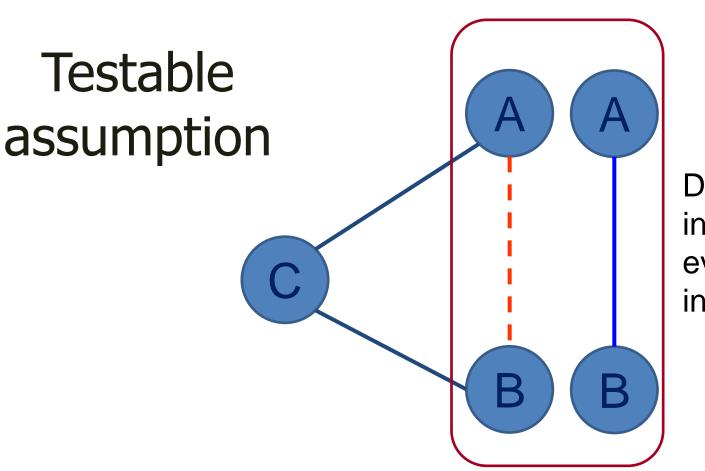
Distribution of mean dose of the active intervention in ten studies



Compare the distribution of important characteristics across treatments (Salanti et al 2009)



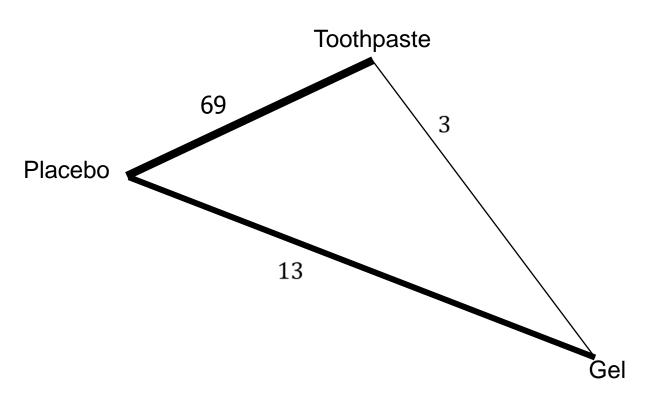
Consistency



Direct and indirect evidence are in agreement

If all three A, B and C are transitive then the loop is consistent

Consistency Equation



Inconsistency Factor

 $\frac{\text{Indirect } SMD_{GvsT}^{ind} = -0.15}{var(SMD_{GvsT}^{ind}) = 0.004}$

Direct $SMD_{GvsT}^{dir} = 0.04$ $var(SMD_{GvsT}^{dir}) = 0.011$

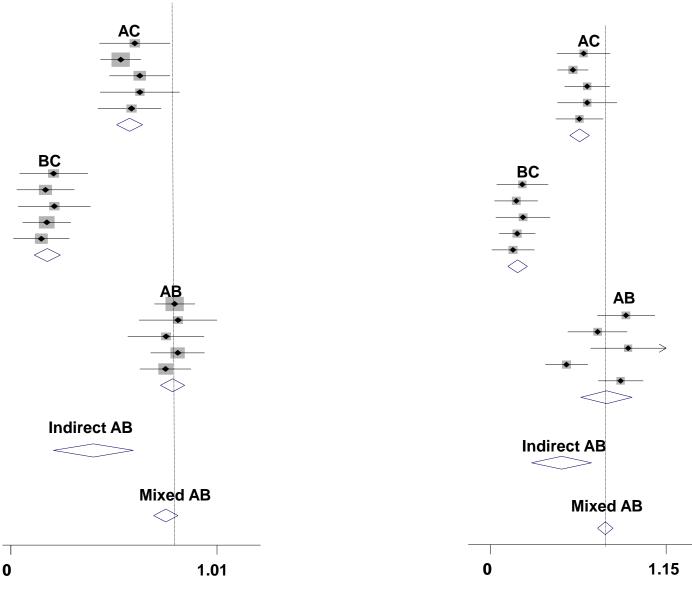
$$IF = |SMD_{GvsT}^{dir} - SMD_{GvsT}^{ind}| = |0.04 - (-0.15)| = 0.19$$

You can do this with any measure... InOR, InRR, RD, mean difference, HR

Consistency and heterogeneity

a) Fixed effects analysis

b) Random effects analysis

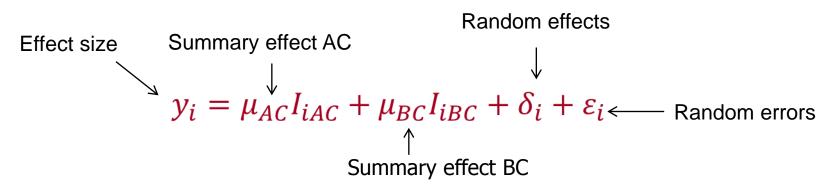


Fit a network meta-analysis model

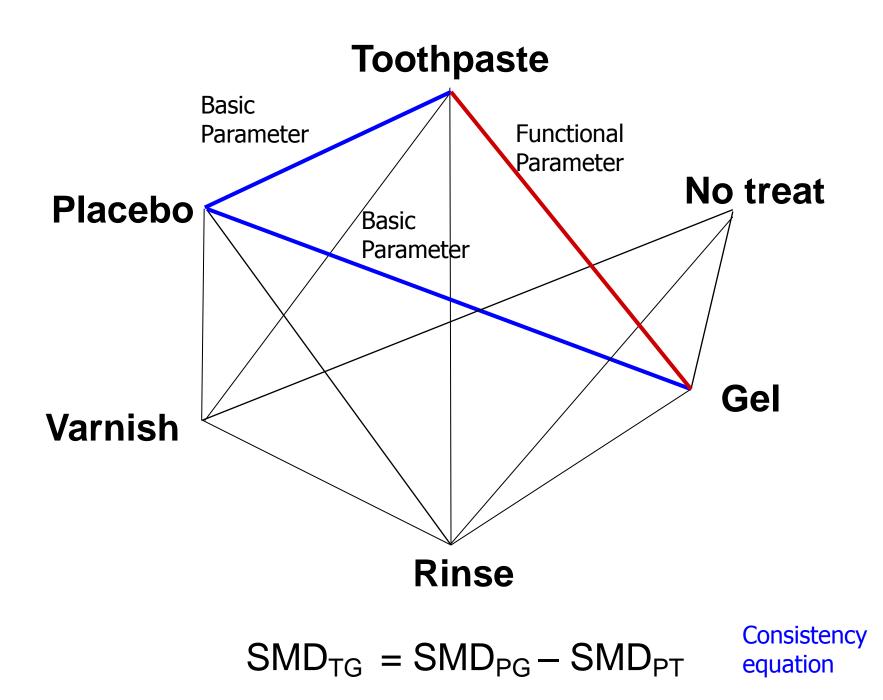
- Meta-analysis is a weighted regression with no covariates
- Network meta-analysis is a weighted regression with dummy variables for the treatments
- You should take into account correlations in multi-arm trials

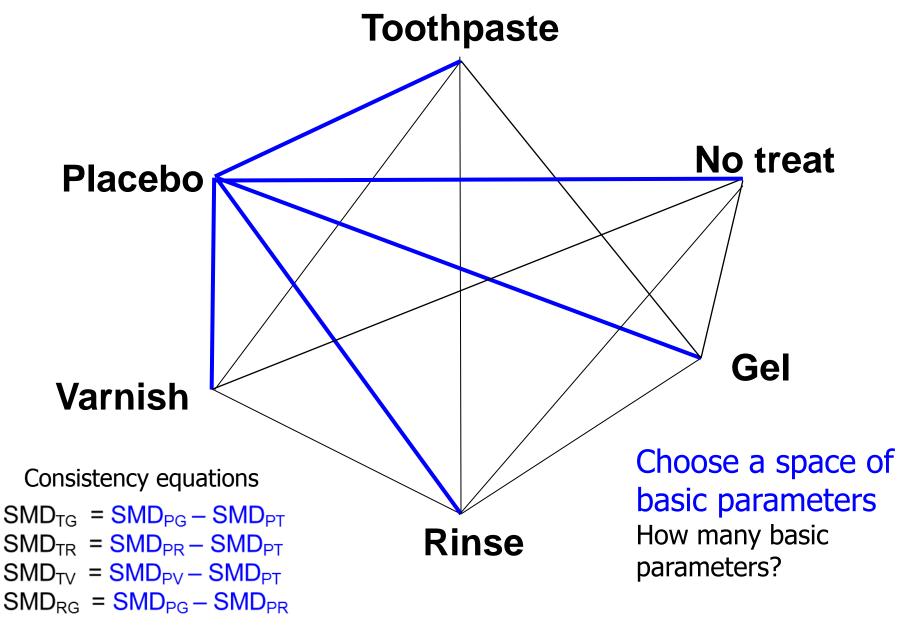
Network and meta-regression

- Meta-regression using the treatments as 'covariates' and without intercept
- With 3 treatments and AC, AB, BC studies, chose C as *reference,* so AC and BC are *basic parameters*



- The AC studies have $(I_{iAC}, I_{iBC}) = (1,0)$, the BC studies $(I_{iAC}, I_{iBC}) = (0,1)[basic]$
- AB studies have $(I_{iAC}, I_{iBC}) = (1, -1)$ [functional] because AB=AC-BC





$y_i = \mu^{PT} T_i + \mu^{PG} G_i + \mu^{PR} R_i + \mu^{PV} V_i + \mu^{PN} N_i$							
Use as 'covariates'							
No. studies	Placebo	Toothpaste	Gel	Rinse	Varnish	NoTreatment	
69	-1	1		0	0	0	
13	-1	0	1	0	0	0	
31	-1	0	0	1	0	0	
3	-1	0	0	0	1	0	
4	0	0	0	-1	0	1	
4	0	0	0	0	-1	1	
9	0	0	-1	0	0	1	
4	0	0	0	-1	1	0	
6	0	-1	0	1	0	0	

 $\mathbf{y} = X(\mu^{PT}, \mu^{PG}, \mu^{PR}, \mu^{PV}, \mu^{PN})' + \boldsymbol{\delta} + \boldsymbol{\varepsilon}$ Vector of Matrix of all Random Design Random observations summary effects matrix errors effects

$$\begin{split} \boldsymbol{\mathcal{E}} \sim N(\boldsymbol{0}, diag(\boldsymbol{v}_i)) & \boldsymbol{\delta} \sim N(\boldsymbol{0}, diag(\tau^2)) \\ \uparrow & \uparrow \\ \text{Variances matrix (for the observed SMD)} & \boldsymbol{\delta} \sim N(\boldsymbol{0}, diag(\tau^2)) \\ & \boldsymbol{\delta} \sim$$

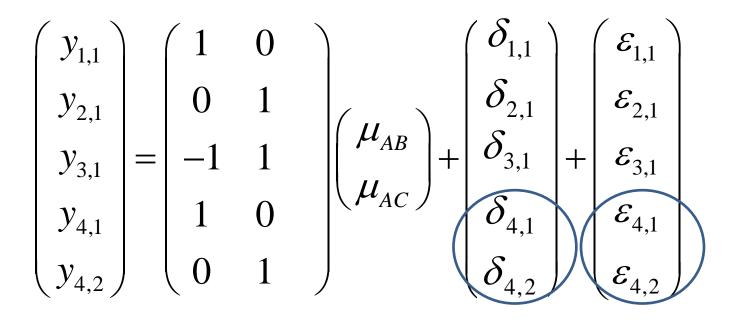
What's the problem with multiarm trials?

- We need to take into account the correlations between the estimates that come from the same study
- A B C $y_{iBC} \varepsilon_{iBC}$ $y_{iAC} \varepsilon_{iAC}$
- The random effects that refer to the same trial <u>are</u> <u>correlated</u> as well
- You have to built in the correlation matrix for the observed effects, **and** the correlation matrix for the random effects

$$\varepsilon \sim N(\mathbf{0}, S)$$
 $\delta \sim N(\mathbf{0}, \Delta)$

Study	No. arms	#	Data	Contrast
i=1	T ₁ =2	1	<i>y</i> _{1,1} , <i>v</i> _{1,1}	AB
i=2	T ₂ =2	1	<i>y</i> _{2,1} , <i>v</i> _{2,1}	AC
i=3	T ₃ =2	1	y _{3,1} , v _{3,1}	BC
i=4	T ₄ =3	2	$\begin{array}{c} y_{4,1}, v_{4,1} \\ y_{4,2}, v_{4,2} \\ \text{COV}(y_{4,1}, y_{4,2}) \end{array}$	AB AC

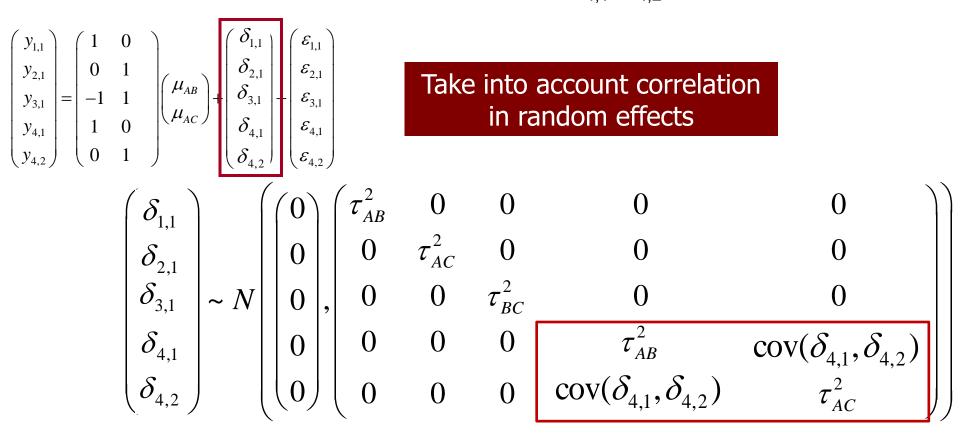
Meta-regression



Study	No. arms	#	Data	Contrast
i=1	T ₁ =2	1	<i>y</i> _{1,1} , <i>v</i> _{1,1}	AB
i=2	T ₂ =2	1	<i>y</i> _{2,1} , <i>v</i> _{2,1}	AC
i=3	T ₃ =2	1	<i>y</i> _{3,1} , <i>v</i> _{3,1}	BC
i=4	T ₄ =3	2	$\begin{array}{c} y_{4,1}, v_{4,1} \\ y_{4,2}, v_{4,2} \\ \text{COV}(y_{4,1}, y_{4,2}) \end{array}$	AB AC

$$\begin{pmatrix} y_{1,1} \\ y_{2,1} \\ y_{3,1} \\ y_{4,2} \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ -1 & 1 \\ 1 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} \mu_{AB} \\ \mu_{AC} \end{pmatrix} + \begin{pmatrix} \delta_{1,1} \\ \delta_{2,1} \\ \delta_{4,1} \\ \delta_{4,1} \\ \delta_{4,2} \end{pmatrix} \begin{pmatrix} \varepsilon_{1,1} \\ \varepsilon_{2,1} \\ \varepsilon_{3,1} \\ \varepsilon_{4,1} \\ \varepsilon_{4,1} \\ \varepsilon_{4,2} \end{pmatrix} \sim N \begin{pmatrix} \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} v_{1,1} & 0 & 0 & 0 & 0 \\ 0 & v_{2,1} & 0 & 0 & 0 \\ 0 & 0 & v_{3,1} & 0 & 0 \\ 0 & 0 & 0 & v_{4,1} & \operatorname{cov}(y_{4,1}, y_{4,2}) \\ 0 & 0 & 0 & \operatorname{cov}(y_{4,1}, y_{4,2}) \end{pmatrix} \end{pmatrix}$$

Study	No. arms	#	Data	Contrast
i=1	T ₁ =2	1	<i>y</i> _{1,1} , <i>v</i> _{1,1}	AB
i=2	T ₂ =2	1	<i>y</i> _{2,1} , <i>v</i> _{2,1}	AC
i=3	T ₃ =2	1	y _{3,1} , v _{3,1}	BC
i=4	T ₄ =3	2	$\begin{array}{c} y_{4,1}, v_{4,1} \\ y_{4,2}, v_{4,2} \\ \text{COV}(y_{4,1}, y_{4,2}) \end{array}$	AB AC



Multivariate meta-analysis

- Studies typically report many outcomes
 Efficacy and acceptability in antidepressants
- Multivariate meta-analysis allows a joint synthesis of the multiple end points
- Different between-treatment contrasts are viewed as different outcomes
- White et al estimate NMA models by expressing them as multivariate randomeffects meta-regressions (mvmeta in STATA)

Data: *n* studies with 2 outcomes Acceptability/AC Efficacy/AB Efficacy/AB Acceptability/AC Study 1: $y_{11}, y_{12}, \begin{pmatrix} s_{11}^2 & S_{112} \\ S_{112} & s_{12}^2 \end{pmatrix}$ Study $i: y_{i1}, y_{i2}, \begin{pmatrix} s_{i1}^2 & S_{i12} \\ S_{i12} & s_{i2}^2 \end{pmatrix}$ Study $n: y_{n1}, y_{n2}, \begin{pmatrix} s_{n1}^2 & S_{n12} \\ S_{n12} & s_{n2}^2 \end{pmatrix}$

Network meta-analysis and multivariate approaches

- We can look at network meta-analysis as either a multivariate meta-regression or a multivariate metaanalysis
- Multivariate meta-regression:
 - extends the meta-regression approach to allow for multi-arm trials
 - dummy 1, -1 and 0 codes for treatments (with a reference in mind)
 - assumes a common heterogeneity variance
- Multivariate meta-analysis:
 - no covariates required
 - Flexible modelling of the between-study variance matrix
 - requires a common reference arm for every study
 - a problem that is surmountable using *data augmentation*

How to fit network meta-analysis?

- R mvmeta, metasem, netmeta
- STATA using metareg (no multi-arm studies)
- STATA mvmeta
- To my knowledge only netmeta in R and mvmeta in STATA model properly the matrix Δ
- Using MCMC (WinBUGS)

Presenting results from network meta-analysis

 With many treatments judgments based on pairwise effect sizes are difficult to make

Antidepressants

Efficacy (response rate)) (95% CI)	Comp	arison 🔲	Acceptability	/ (dropout rat	e) (95% Cl)				
BUP	1.00 (078-1.28)	0.75 (0.55-1.01)	1.06 (0.86-1.32)	0.89 (074-1.08)	0.73 (0.53-1.00)	0.87 (0.58-1.24)	0.87 (0.66-1.14)	0.81 (0.65-1.00)	(0. <u>45-0.86</u>)	1.01 (0.82-1.27)	0.84 (0.68-1.02)
0.98 (0.78–1.23)	сп	0.75 (0.55-1.02)	1.07 (0.86-1.31)	0.90 (073-1.09)	<u>0.73</u> (0.54-0.99)	0.87 (0.60-1.24)	0.87 (0.66-1.15)	0.81 (0.65-1.01)	(0. <u>45-0.84</u>)	1.02 (0.81-1.28)	0.84 (0.67-1.06)
1.09 (0.83-1.43)	1·12 (0·87-1·44)	DUL	<u>1.43</u> (1.09-1.85)	1·19 (0·91-1·57)	0.98 (0.67-1.41)	1.16 (0.77-1.73)	1.16 (0.83-1.61)	1.08 (0.84-1.40)	0.83 (0.57-1.22)	(1·01-1·83)	1.12 (0.84-1.50)
0.82 (0.67-1.01)	0.84 (0.70-1.01)	<u>0.75</u> (0.60-0.93)	ESC	0.84 (0.70-1.01)	0.69 (0.50-0.94)	0.81 (0.55-1.15)	0.81 (0.62-1.07)	0.76 (0.62-0.93)	(0. <u>58</u> (0. <u>43-0.81</u>)	0.95 (0.77-1.19)	(0-64-0-97)
1.08 (0.90-1.29)	1·10 (0·93-1·31)	0.99 (0.79-1.24)	(1·12-1·55)	FLU	0.82 (0.62-1.07)	0.97 (0.69 -1 .32)	0.97 (0.77-1.21)	0.91 (0.79-1.05)	(0. <u>53-0.9</u> 2)	1·14 (0·96–1·36)	0.94 (081-1-09)
1.10 (083-1.47)	1·13 (0·86–1·47)	1.01 (0.74-1.38)	<u>1.35</u> (1.02-1.76)	1.02 (0.81-1.30)	FVX	1.18 (0.76-1.75)	1.18 (0.87-1.61)	1.10 (0.84-1.47)	0.85 (0-57-1-26)	<u>1.38</u> (1.03-1.89)	1·14 (0·86-1·54)
1.07 (0.77-1.48)	1.09 (0.78-1.50)	0.97 (0.69 -1 .38)	1.30 (0.95-1.78)	0.99 (0.74-1.31)	0.97 (0.68 -1 .37)	MIL	0.99 (0.69-1.53)	0.94 (0.68-1.31)	0.72 (0.48-1.10)	1·17 (0·84-1·72)	0.97 (0.69-1.40)
0.79 (0.72-1.00)	0.80 (0.63-1.01)	(0.54-0.94)	0.96 (076-1.19)	0-73 (0-60-0-88)	(0.55-0.92)	0.74 (0.53-1.01)	MIR	0.93 (0.75-1.17)	0.72 (0.51-1.03)	1·17 (0·91-1·51)	0.97 (0.76-1.23)
1.06 (0.87-1.30)	1.08 (0.90-1.30)	0.97 (0.78-1.20)	<u>1·30</u> (1·10-1·53)	0.98 (0.86-1.12)	0.96 (0.76-1.23)	1.00 (0.74-1.33)	<u>1·35</u> (1·11-1·64)	PAR	0.77 (0.56-1.05)	<u>1·25</u> (1·04-1·52)	1.03 (0.86-1.24)
(1.20-2.16)	(1.25-2.14)	(1.05-2.02)	<u>1.95</u> (1.47-2.59)	(1·16-1·90)	(1·03-2·02)	(1·03-2·18)	2:03 (1:52-2:78)	<u>1.50</u> (1.16-1.98)	REB	(1·19-2·24)	1.34 (0.99-1.83)
0.87 (072-1-05)	0.88 (072-1.07)	0.79 (0.62-1.01)	1.06 (0.88–1.27)	<u>0-80</u> (0-69-0-93)	0.79 (0.61-1.01)	0.81 (0.60-1.11)	1·10 (0·90-1·36)	<u>0-82</u> (0-69-0-96)	<u>0.54</u> (0.41-0.71)	SER	0.82 (0.67-1.00)
0.85 (070-1.01)	0.86 (071-1.05)	(0-60-0-99)	1.03 (0.86-1.24)	0- <u>78</u> (0-68-0-90)	<u>0.77</u> (0.59-0.99)	0.79 (0.58-1.08)	1.08 (0.87-1.33)	<u>0.79</u> (0.67-0.94)	(0. <u>40-0.69</u>)	0.98 (0.82-1.16)	VEN

OR>1 means the treatment in top-left is better

Probabilities

- Estimate for each treatment the probability of being the best
- Rankings are contructed by drawing the coefficients a large number of times from their approximate posterior density
- For each draw, the effect sizes are estimated and the largest effect size is noted

12 new generation antidepressants

paroxetine	reboxetine	paroxetine
sertraline	mirtazapine	duloxetine
citalopram escitalopram	fluvoxamine	escitalopram
fluoxetine	citalopram	milnacipran
fluvoxamine	venlafaxine	sertraline
milnacipran	fluoxetine	bupropion
venlafaxine	paroxetine	milnacipran
reboxetine bupropion	duloxetine	sertraline
mirtazapine	escitalopram	bupropion
duloxetine	milnacipran	fluvoxamine

Probability of being the best

0%

7%

0%

26%

0%

0%

1%

11%

0%

0%

54%

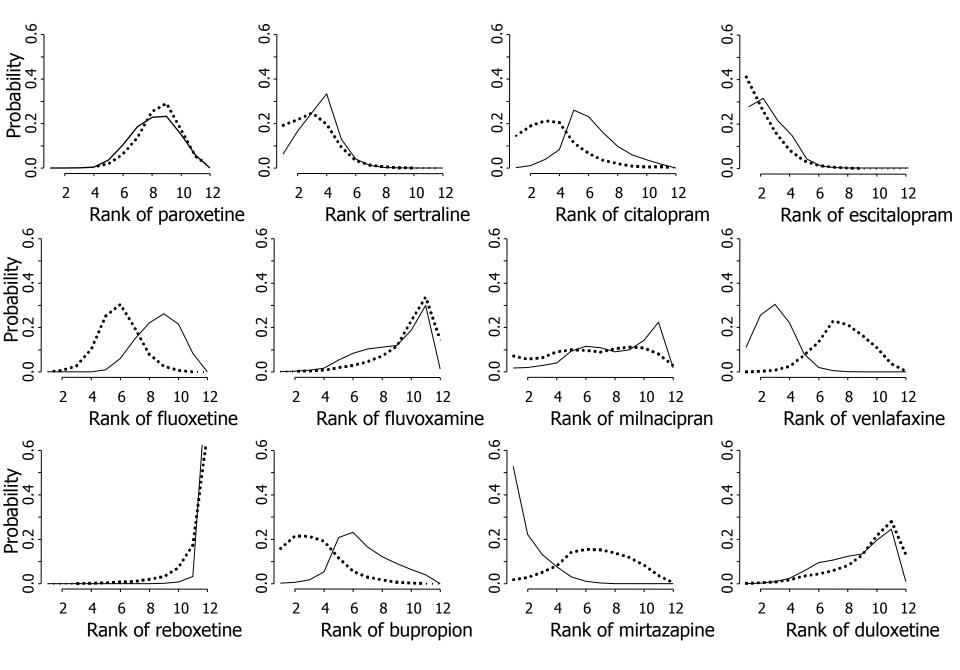
0%

The probability of being the best does not convey the spread of the rank probabilities....

% probability	Α	В	С	D
j=1	0.25	0.50	0.25	0.00

% probability	Α	В	С	D
j=1	0.25	0.50	0.25	0.00
j=2	0.25	0.25	0.50	0.00
<i>j</i> =3	0.25	0.25	0.25	0.25
j=4	0.25	0	0	0.75

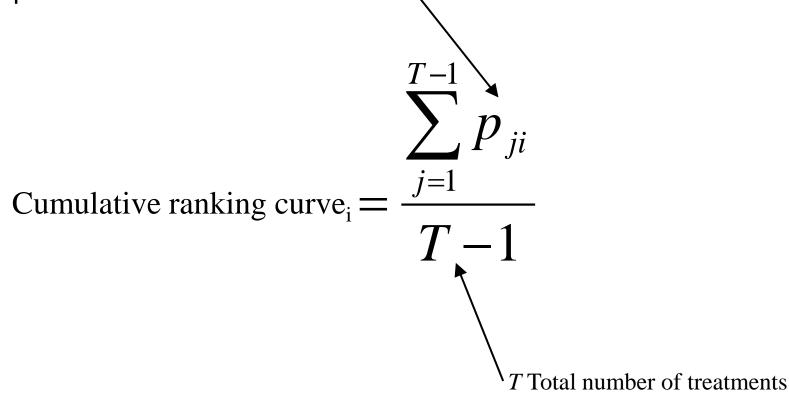
i the treatment *j* the rank



Ranking for efficacy (solid line) and acceptability (dotted line). Ranking: probability of being the best treatment, of being the second best, the third best and so on, among the 12 comparisons).

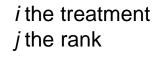
Surface under the cumulative ranking curve

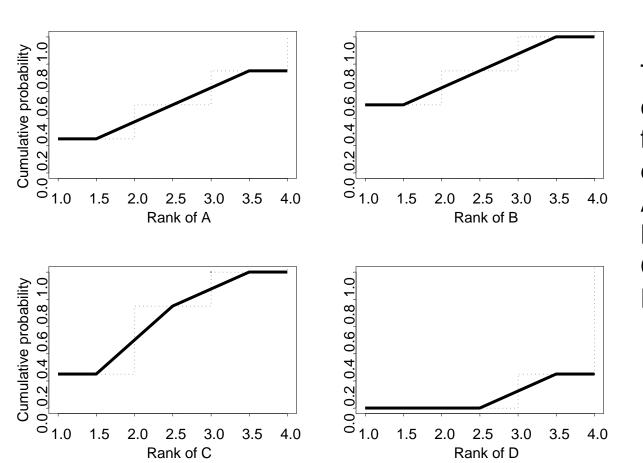
Use <u>posterior probabilities</u> for each treatment to be among the *n*-best options



[J Clin Epidemiol. 2010 Salanti et al]

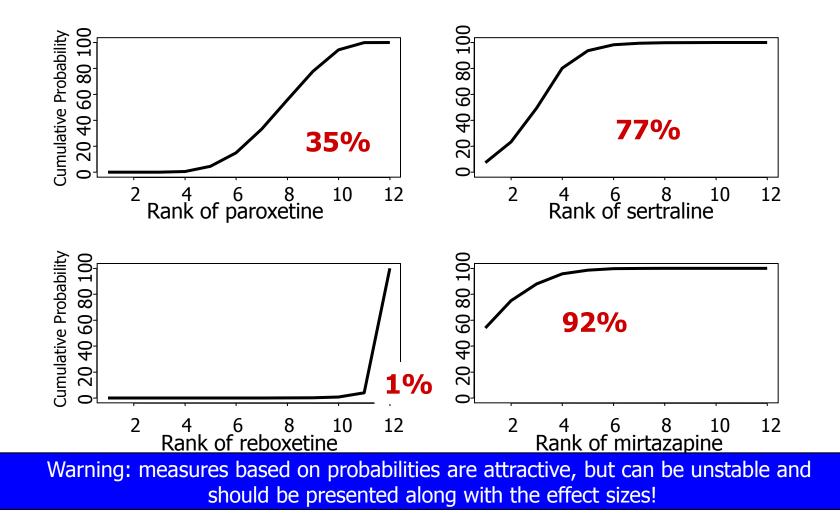
% probability	Α	В	С	D
j=1	0.25	0.50	0.25	0.00
j=2	0.50	0.75	0.75	0.00
j=3	0.75	1.00	1.00	0.25
j=4	1.00	1.00	1.00	1.00





The areas under the cumulative curves for the four treatments of the example above are A=0.5 B=0.75 C=0.67 D=0.08

Surface under the cumulative ranking curve



INCONSISTENCY

Validity of network meta-analysis

- The validity of a network meta-analysis depends on *transitivity* of effect size parameters:
- For any pair A and B, typical (or mean) advantage of A over B = advantage of A over C – advantage of B over C

 In a simple <u>indirect</u> comparison, we <u>cannot</u> test this assumption empirically.

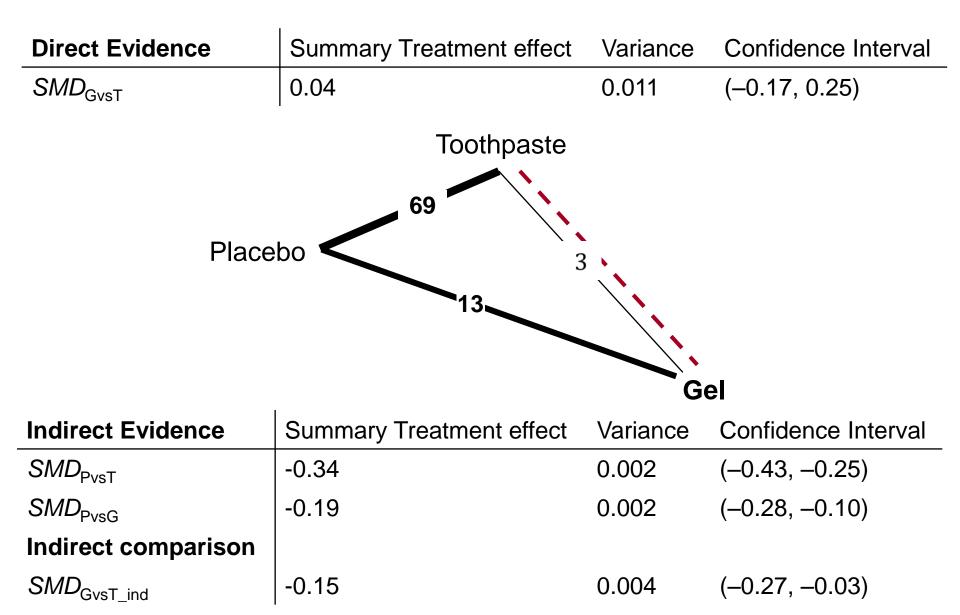
- In a <u>network meta-analysis</u>, we sometimes <u>can</u>.
- We call this looking at *inconsistency*.

Evaluate the assumption of consistency

What is inconsistency?

- Consistency = The data fit together according to the laws of transitivity
- i.e.
 - for each pair of interventions A and B, all sources of evidence about A vs B agree with each other
 - (this means direct evidence (if available) and different routes to indirect evidence)
- *Inconsistency* = Lack of consistency
- Only *closed loops* can tell us about (in)consistency

Example: a simple loop of treatments



How much inconsistency?

- Taking into account the previous evidence,
- the difference between direct and indirect estimates is $IF = \left| SMD_{GvsT}^{dir} - SMD_{GvsT}^{ind} \right| = \left| 0.04 - (-0.15) \right| = 0.19$

 and we add the variances (since the sources of evidence are independent):

Var(difference between direct and indirect) =

$$var(IF) = var(SMD_{GvsT}^{dir}) + var(SMD_{GvsT}^{ind}) = 0.004 + 0.011$$

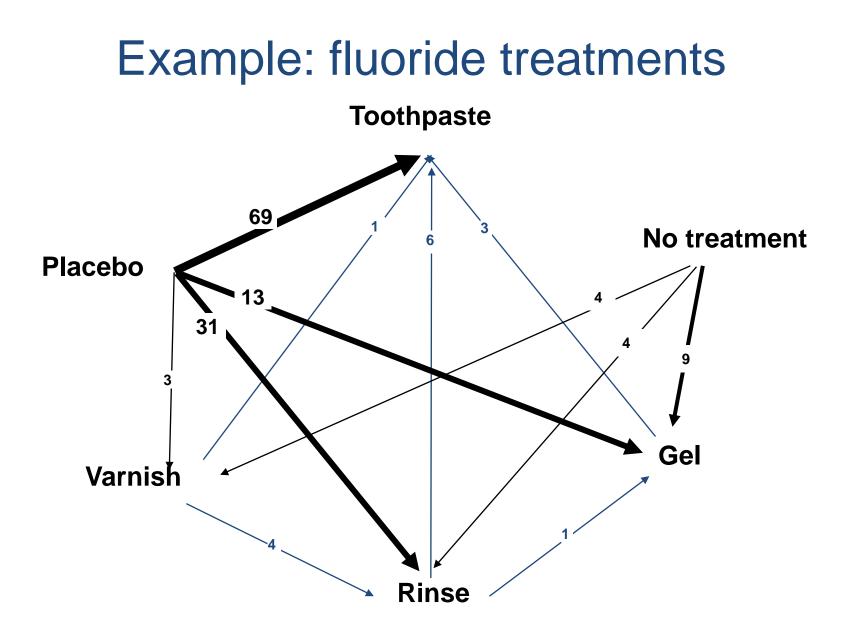
= 0.015

How much inconsistency?

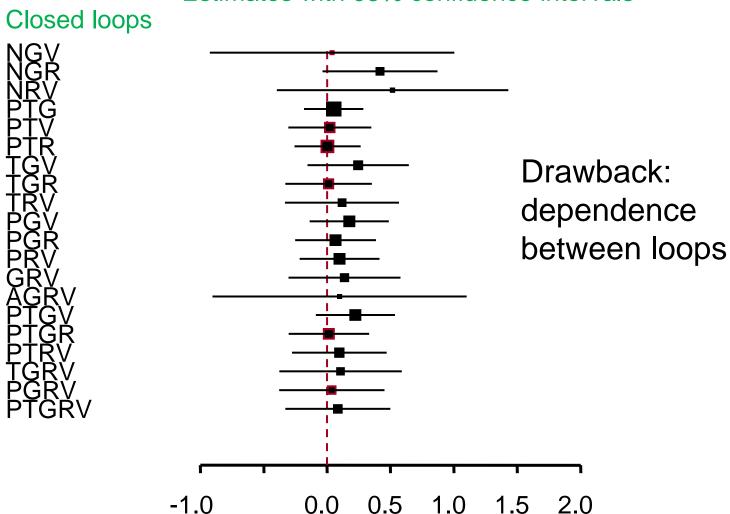
$$Z = \frac{IF}{\sqrt{var(IF)}} \sim N(0,1)$$

95% Confidence Interval for Inconsistency

$$IF \pm 1.96\sqrt{var(IF)} \\ 0.19 \pm 1.96\sqrt{0.015} \\ 0.19 \pm 0.24 \\ (-0.05, 0.43)$$



Evaluation of consistency within closed loops



Estimates with 95% confidence intervals

R routine in www.mtm.uoi.gr/howotodoanmtm.html [Clin Epidemiol 2009, Salanti et al]

Are networks typically inconsistent?

Triangular networks

- Song et al BMJ 2011 found 16/112 (14%) inconsistent triangles
 - The same authors evaluated the assumption of consistency in Cochrane Reviews separately (Xiong et *al* JCE 2013) and found 16/94 (17%) triangles inconsistent

Complex networks

- Veroniki et al (IJE 2013) published network meta-analyses with binary data that involve at least 4 treatments and at least one closed loop
 - so far 40 networks, 303 loops
 - Inconsistency was detected in between 2% and 10% of the tested loops, depending on the effect measure and heterogeneity estimation method
 - About one eighth of the networks was found to be inconsistent.

Approaches for exploring inconsistency

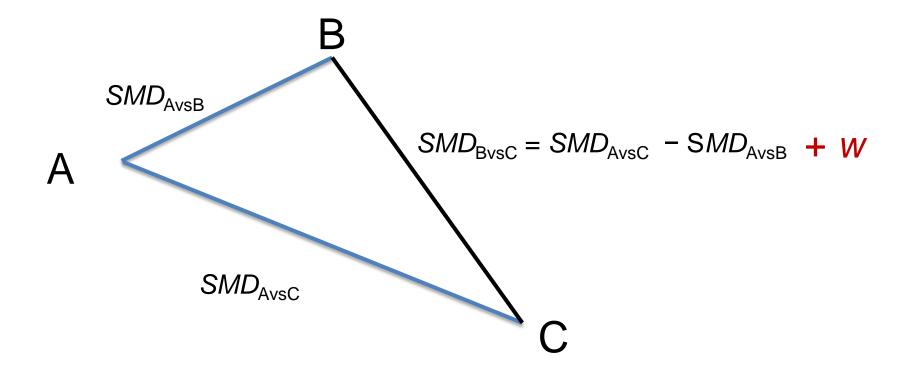
Evaluation of local inconsistency

- Loop-Specific : examine each closed loop separately
- Node-splitting (Dias et *al* Stat Med 2010)

Evaluation of global inconsistency

- Use a network meta-analysis model that allows for inconsistency (Lu & Ades JASA 2005)
- Compare model fit between consistency and inconsistency models
- Apply a 'design by treatment' interaction model (White et al RSM 2012, Higgins et al RSM 2012)

Inconsistency models: introduction



Model for consistency

Modelled log odds ratios						
(basic parameters μ_{AB} and μ_{AC});						
∂_i is th	e neteroge	eneity random	effect			
Trial	A	В	С			
AB	ref	$\mu_{AB} + \delta_i$				
AC	ref		$\mu_{AC} + \delta_i$			
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$			

Model for inconsistency Lu and Ades model

A

Modelled log odds ratios							
	(basic parameters μ_{AB} and μ_{AC});						
δ_i is th	e heteroge	eneity random	effect				
Trial	A	В	С				
AB	ref	$\mu_{AB} + \delta_i$					
AC	ref		$\mu_{AC} + \delta_i$				
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i + w$				

Model for consistency with a three-arm trial

Modelled log odds ratios							
(ba	(basic parameters μ_{AB} and μ_{AC});						
δ_i is th	e heteroge	eneity random	effect				
Trial	A	В	С				
ABC	ref	$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$				
AB	ref	$\mu_{AB} + \delta_i$					
AC	ref		$\mu_{AC} + \delta_i$				
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$				

Issues with the Lu and Ades model

- In the presence of multi-arm trials, the Lu and Ades inconsistency model is not uniquely defined
- Problems arise because multi-arm trials *must* be consistent, so a network with multi-arm trials will have a mixture of **consistent** and **inconsistent** loops

Lu and Ades model

Modelled log odds ratios(basic parameters μ_{AB} and μ_{AC}); δ_i is the heterogeneity random effect				
Trial	A	В	С	
ABC	ref	$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$	
AB	ref	$\mu_{AB} + \delta_i$		
AC	ref		$\mu_{AC} + \delta_i$	
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i + w$	

Design by treatment interaction model

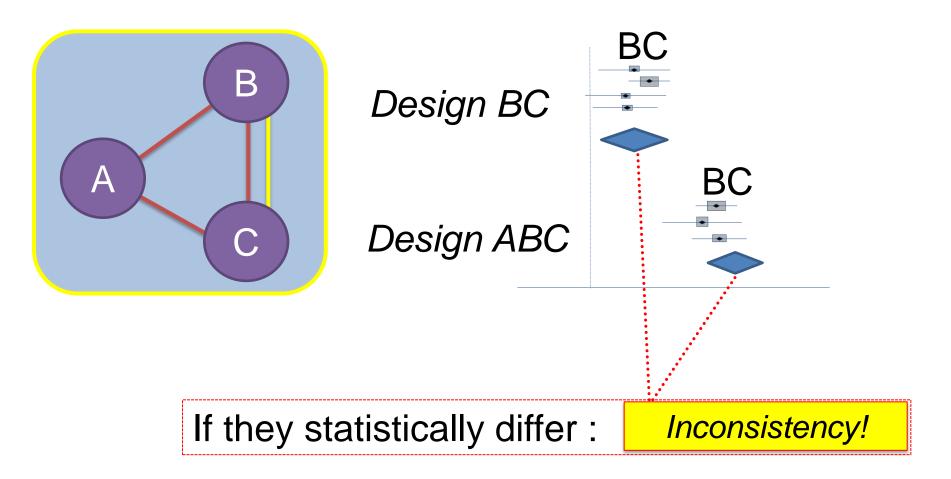
- A model that is completely general is one that allows for all types of inconsistency
 - inconsistency within loops made up of different trials
 - inconsistency between two-arm and three-arm trials
 - and beyond...
- Such a model has been termed a design-by-treatment interaction model

Forms of Inconsistency

Loop Inconsistency BC B Direct BC AC A Indirect BC С If they statistically differ : **Inconsistency!**

Forms of Inconsistency

Design Inconsistency



Design-by-treatment interaction model

Modelled log odds ratios(basic parameters μ_{AB} and μ_{AC}); δ_i is the heterogeneity random effect				
Trial	A	В	С	
ABC	ref	$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$	
AB	ref	$\mu_{AB} + \delta_i + w_{AB}$		
AC	ref		$\mu_{AC} + \delta_i + w_{AC}$	
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i + w_{BC}$	

[Higgins et al RSM 2012], [White et al RSM 2012]

Lu and Ades model for inconsistency with a three-arm trial

Modelled log odds ratios(basic parameters μ_{AB} and μ_{AC}); δ_i is the heterogeneity random effect				
Trial	A	В	С	
ABC	ref	$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i$	
AB	ref	$\mu_{AB} + \delta_i$		
AC	ref		$\mu_{AC} + \delta_i$	
BC		$\mu_{AB} + \delta_i$	$\mu_{AC} + \delta_i + w$	

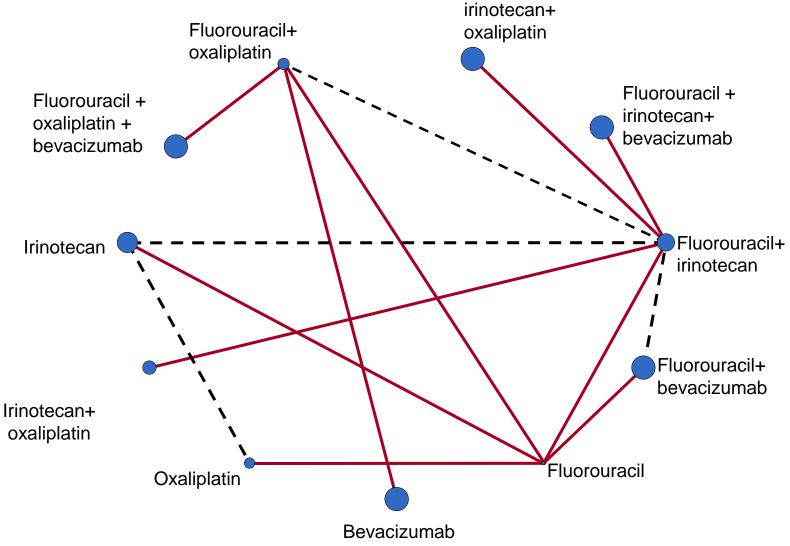
Modelling the *w* parameters

 When we have several inconsistency (w) parameters, we could let them have a random-effects distribution across comparisons

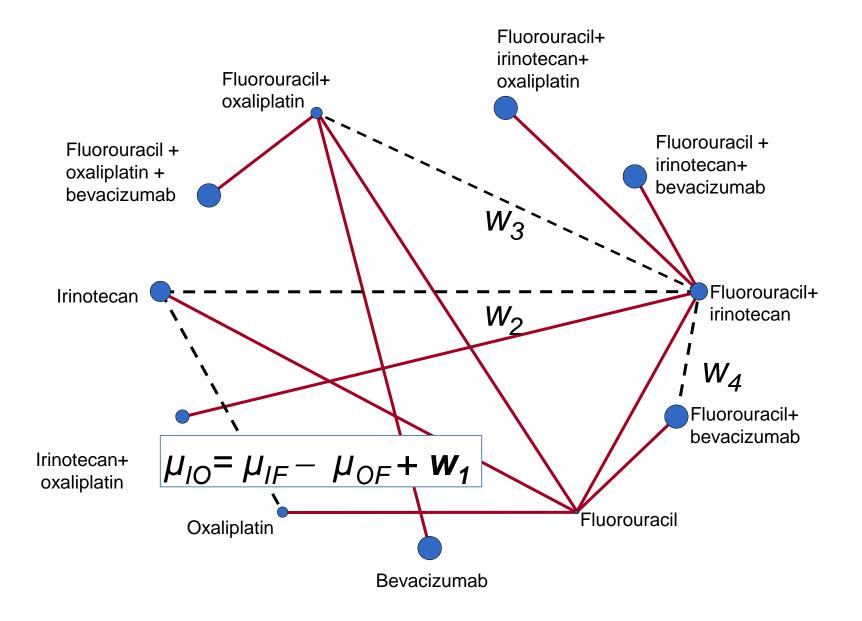
 $W_j \sim N(0,\sigma^2)$

- Comparing σ² with τ² (heterogeneity) allows us to assess the magnitude of the inconsistency
- I prefer to use fixed effects for the *w* parameters
 - can interpret them individually
 - and it's easier to fit the model using Stata

Example: Survival with 11 chemotherapy regimens in colorectal cancer



Lu and Ades model for colorectal cancer



Results: colorectal cancer network

• $w_1 = -0.08$, $w_2 = -0.07$, $w_3 = -0.06$, $w_4 = -0.03$

- No loop is remarkably inconsistent

- σ² = 0.11 (SD 0.04), τ²=0.19 (SD 0.18)
- $P(\sigma^2 > \tau^2) = 0.41$
 - No important changes in posterior HRs or fit

What if we find inconsistency?

- Try to explain inconsistency!
- Use network meta-regression
- Might consider
 - presenting results from the inconsistency model
 - presenting a variety of separate direct, indirect and mixed comparisons
- Be careful! Selective inclusion of evidence pieces might lead to bias

Comparison of assumptions (random effects models)

Meta-analysis

Similarity of participants, interventions and outcomes

Appropriate modelling of study data (within-study variances often assumed known, uncorrelated with effects)

Normal distribution for random effects

Possibly covariates to explain heterogeneity

Network meta-analysis

Similarity of participants, outcomes; 'random selection' of interventions

Appropriate modelling of study data (within-study variances often assumed known, uncorrelated with effects)

Normal distribution for random effects

Possibly covariates to explain heterogeneity and/or inconsistency

Possible assumptions about different τ values for different comparisons

Possible extra parameters to allow for inconsistency across comparisons

Hands on

www.mtm.uoi.gr



September 14, 2011

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CONTACT US	

Multiple-Treatments Meta-analysis (MTM)

Meta-analysis is the statistical technique used to synthesize evidence from experiments addressing the same research question. It is often
used to combine data from clinical trials regarding the relative effectiveness of two interventions in order, for example, to infer about
whether antihypertensives A and B are equally effective in lowering blood pressure.
The main drawback of the current state of the art is that meta-analysis focuses on comparing only two alternatives. However, clinicians
and patients need to know the relative ranking of a set of alternative options and not only whether option A is better than B.
The statistical methodology applied to synthesize information over a network of comparisons involving all alternative treatment options
for the same condition is called Multiple-Treatments Meta-Analysis.

his site provides

- an introduction to statistical and methodological issues related to MTM
- links to training material
- support to statisticians with the analysis of networks of interventions

Hands on

- <u>www.mtm.uoi.gr</u>
- Go to 'how to do an MTM' tab
- Use R routine mtmnetwork.plot to plot a network
- Use the R routine ifplot.fun to plot inconsistency in all closed loops
- In WinBUGS: read the description of models (e.g. <u>www.mtm.uoi.gr/3.binarymodeldescription.pdf</u>) download the data and the WinBUGS code
- Use the R routine sucraplot.fun to get rankograms and SUCRA
- Go to 'STATA routines for Network Meta-Analysis' tab for an implementation of network meta-analysis



Research Synthesis Methods

The official journal of the Society for Research Synthesis Methodology <u>http://www.srsm.org/</u>

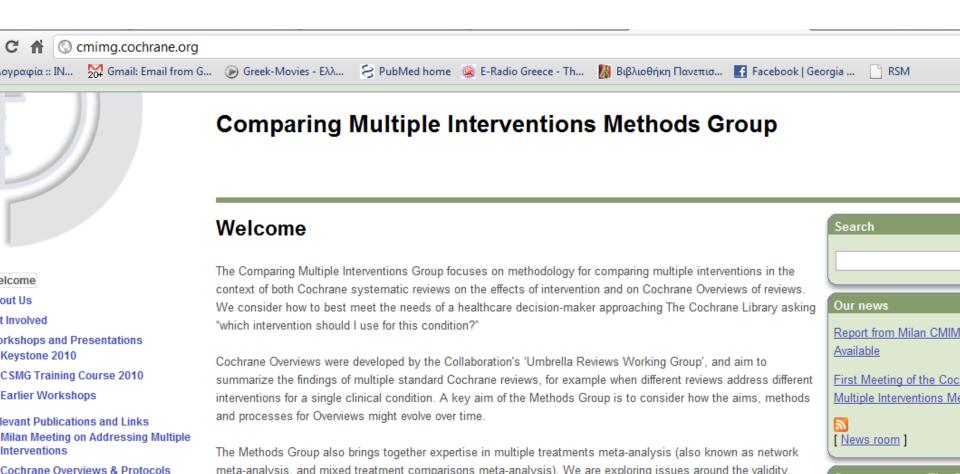
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A special issue for Network Meta-analysis published in 2012

The Cochrane Collaboration

A new methods group has been recently established to support reviews that aim to compare multiple interventions <u>cmimg.cochrane.org</u>



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Thank you! Questions?